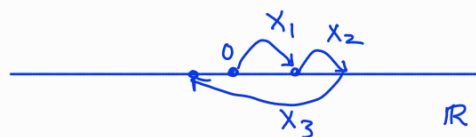


- X is a r.v s.t. $\mathbb{E} e^{\lambda X} < \infty \quad \forall \lambda \in \mathbb{R}$.
 $\Rightarrow \mathbb{E} e^{\lambda X} \leq e^{\lambda^2 \sigma^2 / 2} \quad \forall |\lambda| \leq \lambda_0$
 $\Rightarrow X$ is sub exponential.



- X_n iid X and $S_n = \sum_{i=1}^n X_i$.



$$\mathbb{P}(S_n - n\mu \geq c_n) \leq \exp\left(-\frac{kc_n^2}{n}\right)$$

- Typical value $\approx \sqrt{n}$ $S_n = O_p(\sqrt{n})$
 $\Leftrightarrow \forall \varepsilon > 0, \exists C$ s.t.
 $\mathbb{P}(S_n \geq C\sqrt{n}) \leq \varepsilon.$

- $\mathbb{P}\left(\frac{S_n}{n} - \mu \geq c\right) \leq \exp(-k_c n) \rightarrow \text{Prove SLLN!}$
 $\downarrow$
 Largest k_c possible $\rightarrow$ [Exc: Use Borel Cantelli Lemma I]

- $\checkmark$ $\text{poly}(n) \exp(-k_c n) \leq \mathbb{P}\left(\frac{S_n}{n} - \mu \geq c\right) \leq \text{poly}(n) \exp(-k_c n)$
 $\rightarrow$ Exact rate of decay. $\checkmark$ Large Deviation Event.

★ Cramer - Chernoff bound: X_n iid X , $\mathbb{E} e^{\lambda X} < \infty \quad \forall \lambda \in \mathbb{R}$, $\mu = \mathbb{E} X$.

$$\begin{aligned} \mathbb{P}\left(\sum_{i=1}^n X_i \geq an\right) &\begin{cases} \rightarrow \text{Should decay exponentially for } a \geq \underline{\mathbb{E} X} \\ \rightarrow 1 - e^{-n} \text{ if } a < \mathbb{E} X \end{cases} \\ &= \mathbb{P}\left(e^{\lambda \sum_{i=1}^n X_i} \geq e^{\lambda an}\right) \quad \forall \lambda > 0 \\ &\leq e^{-\lambda an} \mathbb{E} e^{\lambda \sum_{i=1}^n X_i} = e^{-\lambda an} (\mathbb{E} e^{\lambda X})^n, \quad \text{let } \psi_X(\lambda) = \log \mathbb{E} e^{\lambda X} \\ &= \exp\left(-\lambda an + n \psi_X(\lambda)\right) = \exp\left(-n\left(\lambda a - \psi_X(\lambda)\right)\right) \end{aligned}$$

$$\therefore \mathbb{P}\left(\sum_{i=1}^n X_i \geq an\right) \leq \exp\left(-n \sup_{\lambda \geq 0} (\lambda a - \psi_X(\lambda))\right)$$

Observations: 1. $\psi_X(\lambda)$ - strictly convex, so long as $\text{Var}(X) > 0$.

Assume

$$\psi \in C^2$$

$$\begin{aligned} \psi_X''(\lambda) &= \mathbb{E} \left[X^2 \frac{e^{\lambda X}}{\mathbb{E} e^{\lambda X}} \right] - \left(\mathbb{E} \left[X \frac{e^{\lambda X}}{\mathbb{E} e^{\lambda X}} \right] \right)^2 \\ &= \text{Var}_{\mu}(X) > 0 \end{aligned}$$

$$2. \sup_{\lambda \in \mathbb{R}} (\lambda a - \psi(\lambda)) = \psi^*(a) \leftarrow \text{Convex dual of } \psi.$$

$$\text{Max. at } \begin{cases} \psi'(\lambda^*) = a > \mathbb{E} X \\ \psi'(0) = \mathbb{E} X \end{cases}$$

$$\psi'(\lambda) = \mathbb{E} \left[X \frac{e^{\lambda X}}{\mathbb{E} e^{\lambda X}} \right]$$

Since ψ is strictly convex, ψ' is strictly $\uparrow$.

$$\therefore \lambda^* > 0.$$

$$\circ \circ \quad \mathbb{P} \left(\frac{S_n}{n} \geq a \right) \leq \exp(-n \psi^*(a)) \quad \text{Chernoff Cramer bound}$$

Right Tail Chernoff bd $\nearrow$ $a > \mathbb{E} X$

egs: 1. $X \sim N(0,1), \psi_X(\lambda) = \frac{\lambda^2}{2},$

$$\psi_X^*(a) = \sup_{\lambda \in \mathbb{R}} \left\{ \lambda a - \frac{\lambda^2}{2} \right\} \xrightarrow{\lambda=a} = \frac{a^2}{2}$$

$$\mathbb{P} \left(\frac{S_n}{n} \geq a \right) \leq \exp \left(-n \frac{a^2}{2} \right)$$

Left Tail:

$$\begin{aligned} &\mathbb{P} \left(\frac{S_n}{n} \leq a \right) \text{ for } a < \mathbb{E} X \\ &= \mathbb{P} \left(e^{\lambda S_n} \geq e^{\lambda a n} \right) \text{ for } \lambda < 0 \\ &= \exp \left(-n \sup_{\lambda < 0} (\lambda a - \psi(\lambda)) \right), \quad a < \mathbb{E} X \\ &= \exp \left(-n \sup_{\lambda \in \mathbb{R}} (\lambda a - \psi(\lambda)) \right) = \exp \left(-n \psi_X^*(a) \right) \end{aligned}$$

✓

$$\textcircled{\star} \quad \begin{aligned} \mathbb{P}(S_n \geq an) &\leq \exp(-n \psi^*(a)) \quad \text{if } a > \mathbb{E}X \\ \mathbb{P}(S_n \leq an) &\leq \exp(-n \psi^*(a)) \quad \text{if } a < \mathbb{E}X \end{aligned}$$

Automatically gives
us SLLN

0th result of
Large Deviation Theory.

$$\rightarrow \lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{P}(S_n \geq an) = -\psi^*(a)$$

★ Thm (Cramer's theorem): $\mathbb{E} e^{sX} < \infty \quad \forall |s| \leq s_0$, then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{P}(S_n \geq an) = -\psi^*(a) \quad \text{Large dev. event} \quad \text{LD rate f.} \quad \forall a > \mathbb{E}X$$

$$\lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{P}(S_n \leq an) = -\psi^*(a) \quad \forall a < \mathbb{E}X.$$

eg: 1. If $X \sim N(0,1)$,

$$\bullet S_n = \sum_{i=1}^n X_i, \quad X_i \text{ iid } N(0,1)$$

$$\left(\frac{1}{a} - \frac{1}{a^3}\right) e^{-\frac{a^2}{2}} \leq \mathbb{P}(X \geq a) \leq \frac{1}{a} e^{-\frac{a^2}{2}}$$

$$\mathbb{P}(S_n \geq na) = \mathbb{P}\left(\frac{S_n}{\sqrt{n}} \geq \sqrt{n}a\right), \quad \frac{S_n}{\sqrt{n}} \sim N(0,1)$$

$$\bullet \left(\frac{1}{\sqrt{n}a} - \frac{1}{(\sqrt{n}a)^3}\right) e^{-\frac{na^2}{2}} \leq \cdot \leq \frac{1}{\sqrt{n}a} e^{-\frac{na^2}{2}}$$

★ Ex: $\text{Ber}(p)$, $\text{Poisson}(\lambda)$, $\text{Exp}(\lambda)$,

For these, try to the Cramer-Chernoff bds.

$$\begin{aligned}
 \star \quad \mathbb{P}\left(\frac{S_n}{n} \in [a, \infty)\right) &\rightsquigarrow \mathbb{P}\left(\frac{S_n}{n} \in [a, b)\right) \\
 &= \mathbb{P}\left(\frac{S_n}{n} \in [a, \infty)\right) - \mathbb{P}\left(\frac{S_n}{n} \in [b, \infty)\right) \\
 &\approx (\cdot) e^{-n\psi^*(a)} - (\cdot) e^{-n\psi^*(b)} \\
 \psi^*(a) < \psi^*(b) &\longrightarrow \text{Rate of decay} = -\psi^*(a)
 \end{aligned}$$

$$\bullet \frac{1}{n} \log \left(\mathbb{P}\left(\frac{S_n}{n} \in A\right) \right) = - \inf_{x \in A} \psi^*(x) \xrightarrow{\text{sufficiently regular}} \text{Large Deviation Principle.}$$

⊛ Concentration Inequalities for $X_1 + \dots + X_n$:

1. Chebychev's Ineq: Variance bounds $\Rightarrow$ Conc. ineq. ✓
2. Exponential Conc. Ineq. $\begin{cases} \nearrow \text{Hoeffding} \xrightarrow{\text{green}} \text{Depends on } \text{sg norm} \\ \searrow \text{Bernstein} \xrightarrow{\text{green}} \text{Depends on variance} \end{cases}$
3. Cramer Chernoff method $\rightarrow$ More direct exp. conc. ✓

$$\text{II. } \mathbb{P}\left(|f(X_1, X_2, \dots) - \mathbb{E}f(X_1, \dots)| \geq c_n\right) = o(1)$$

$$\text{eg. } f = \frac{X_1 + \dots + X_n}{n} \checkmark, \quad \frac{X_1^2 + \dots + X_n^2}{n} \checkmark, \dots \text{ what others}$$

- “ $f(X_1, \dots, X_n)$ conc around its mean if
- $\rightarrow$ i) $X_1, \dots, X_n$ are weakly ~~in~~ dependent
 - ii) f depends weakly on each of $X_1, \dots, X_n$.”

2. Martingales:

Defⁿ: $X_0, X_1, X_2, \dots$ is a collection of random variables. Then a collection $\{M_t\}_{t \in \mathbb{N} \cup \{0\}}$ is called a Martingale w.r.t

$X_0, X_1, X_2, \dots$ if

- ✓ i) $X_0, \dots, X_t$ completely determines M_t ← Amount of money gained
- ✓ ii) $\mathbb{E}|M_t| < \infty$ and
- iii) $\mathbb{E}[M_{t+1} | X_0, X_1, \dots, X_t] = M_t$

eg. 1. $X_1, \dots$ iid X , $\mathbb{E}X = \mu$.

Then $M_n = (X_1 + \dots + X_n) - n\mu$

$$\begin{aligned} \mathbb{E}[M_{t+1} | X_1, \dots, X_t] &= \mathbb{E}[X_1 + \dots + X_t + \underbrace{X_{t+1}}_{-(t+1)\mu} | X_1 + \dots + X_t] \\ &= (X_1 + \dots + X_t) + \mathbb{E}[X_{t+1}] - (t+1)\mu \\ &= M_t \end{aligned}$$

∴ This a martingale.

• Property: $\mathbb{E}M_t = \mathbb{E}M_0$, t a fixed time.

Proof: $\mathbb{E}M_{t+1} = \mathbb{E}M_t = \dots$ (use Tower law).

2. $M_n = S_n^2 - n\sigma^2$ if $\text{Var } X = \sigma^2 < \infty \rightarrow$ Also defines a martingale.

Defⁿ: 1. Supermartingale : $\mathbb{E}[M_{t+1} | \overbrace{X_0, \dots, X_t}^{\mathcal{F}_t}] \leq M_t$
 Submartingale : $\mathbb{E}[M_{t+1} | \mathcal{F}_t] \geq M_t$

Sup : $\mathbb{E}M_0 \geq \mathbb{E}M_1 \geq \dots$

Sub : $\mathbb{E}M_0 \leq \dots$

Defⁿ: 1. Stopping time : A stopping time $X_0, X_1, X_2, \dots$ is a random variable τ s.t. $\{\tau \leq t\}$ is completely determined by $X_0, \dots, X_t$.

2. Predictable process: $(A_t)_{t \geq 0}$ s.t. A_t is compl. det. by $X_0, X_1, \dots, X_{t-1}$.

eg. Let Y be a rv depending upon $X_0, X_1, X_2, \dots$, assume $E|Y| < \infty$.
Then, $M_t = E[Y | X_0, \dots, X_t]$ is a martingale.

$$E[M_t | X_0, \dots, X_{t-1}] = E[E[Y | X_0, \dots, X_t] | X_0, \dots, X_{t-1}]$$

Tower Law

$$= E[Y | X_0, \dots, X_{t-1}] = M_{t-1}.$$

(Doob's Martingale OR Information Exposure Martingale)

* Property: $M_0, M_1, \dots$ are martingales wrt $X_0, X_1, \dots$, then

• $Cov(M_k - M_\ell, M_s - M_t) = 0$ if



so long as $[k, \ell] \cap [s, t] = \emptyset$. [Exercise]

Idea: $(-f(X_1, X_2, \dots, X_n) + E f(X_1, X_2, \dots, X_n))$

$$= (E f(\underline{x}) - E[f(\underline{x}) | X_1]) + (E[f(\underline{x}) | X_1] - E[f(\underline{x}) | X_1, X_2])$$

$$+ \dots + (E[f(\underline{x}) | X_1, \dots, X_{n-1}] - f(\underline{x}))$$

• $f - Ef =$ Sum of uncorrelated random variables.
(using Doob's Martingale).

• $E e^{X_1 + X_2} = E e^{X_1} E e^{X_2}$

→ We would get around these difficulties using monotonicity of Submartingale.

Thm: (Optional Stopping Theorem):

- (M_t) is a submartingale. Then

$$\mathbb{E} M_T \geq \mathbb{E} M_0$$

Requires some cond^{ns}

- i) T is bounded ✓
- ii) .
- iii) .
- iv) $M_t \geq 0$

Real qⁿ:

$$\mathbb{E} M_T \geq \mathbb{E} M_0 ? \quad \forall \text{ sub } M$$

Let (M_t) be a M , then M_t & $-M_t$ are both sub M

$$\mathbb{E}[-M_T] \geq \mathbb{E}[-M_0] \quad , \quad \mathbb{E} M_T \geq \mathbb{E} M_0$$

$$\Rightarrow \boxed{\mathbb{E} M_T = \mathbb{E} M_0}$$

★

Week 1

Week 2

Week 3 -

Week 4 - 9/06 , 12/06

$$\left[\begin{array}{cc} -16/06 & 19/06 \\ - & \end{array} \right] \rightarrow \text{No classes}$$

Weeks 5 - ~~30/06~~

- | -
 . | -

Week 8 -